

Robustness of groups and trajectories in Nagin's finite mixture model

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joint work with
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Outline

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General description of Nagin's model

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This is still an inter-individual model, but unlike other classical models such as standard growth curve models, it allows the existence of subpopulations with completely different behaviors.

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Aim of the analysis: Find r groups of trajectories of a given kind (for instance polynomials of degree 4, $P(t) = \beta_0 + \beta_1 t + \beta_2 t^2 + \beta_3 t^3 + \beta_4 t^4$.)

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- finite : sums across a finite number of groups
- mixture : population composed of a mixture of unobserved groups



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Software:

SAS-based Proc Traj procedure

by Bobby L. Jones (Carnegie Mellon University).

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Rule:

The bigger the BIC, the better the model!

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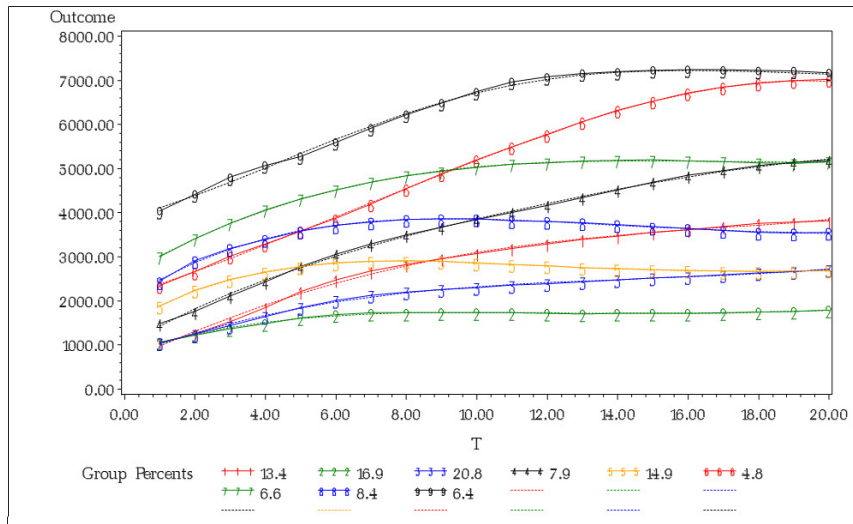
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Bigger groups have on average larger probability estimates.

To be classified into a small group, an individual really needs to be strongly consistent with it.

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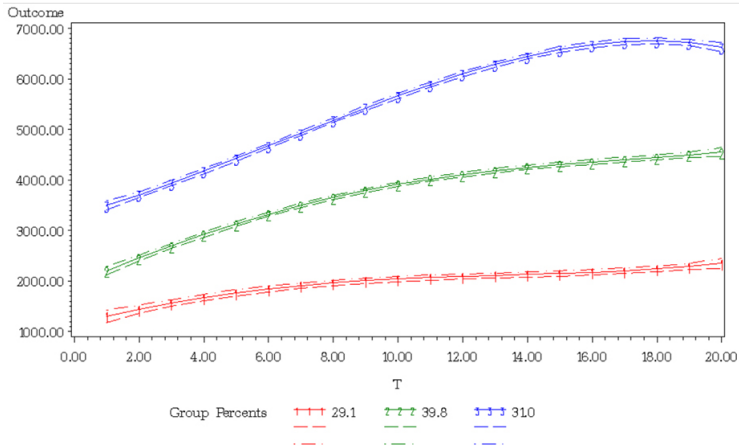


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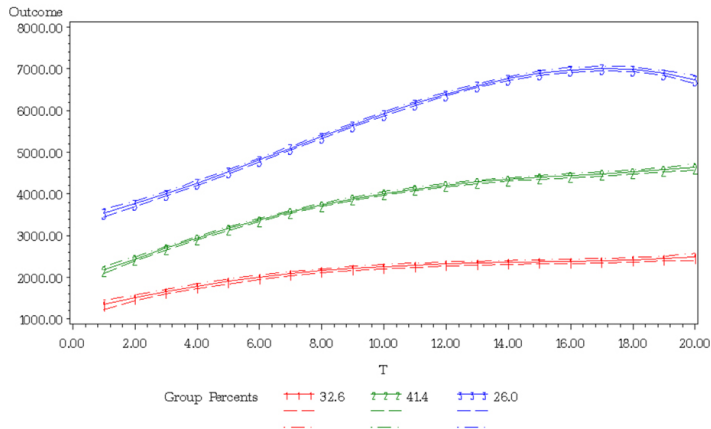
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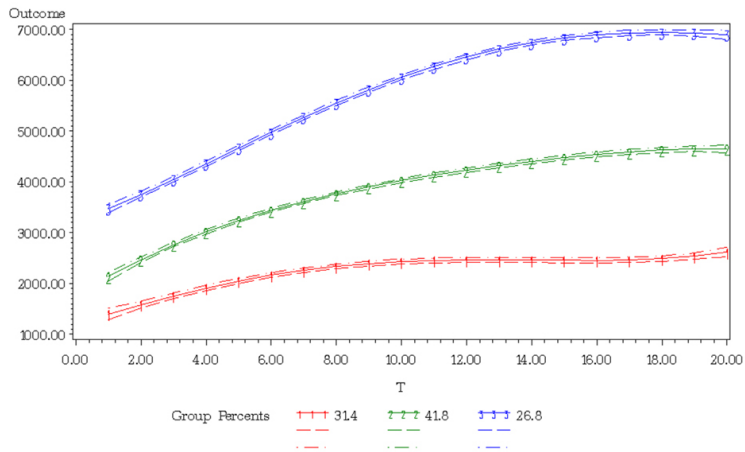
Result for 3 groups : workers beginning their career in 1982



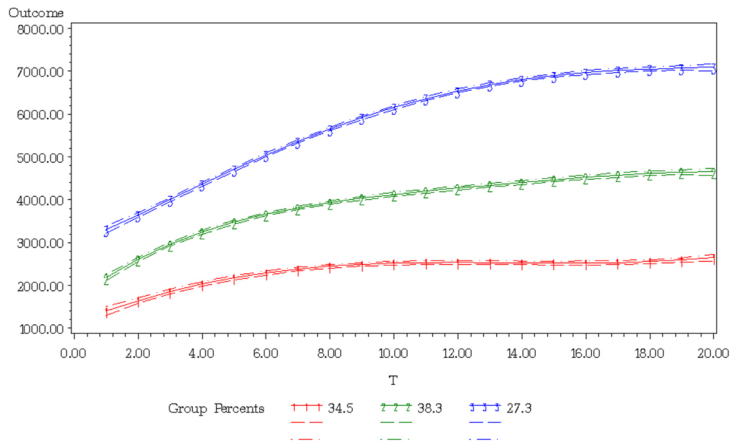
Result for 3 groups : workers beginning their career in 1983



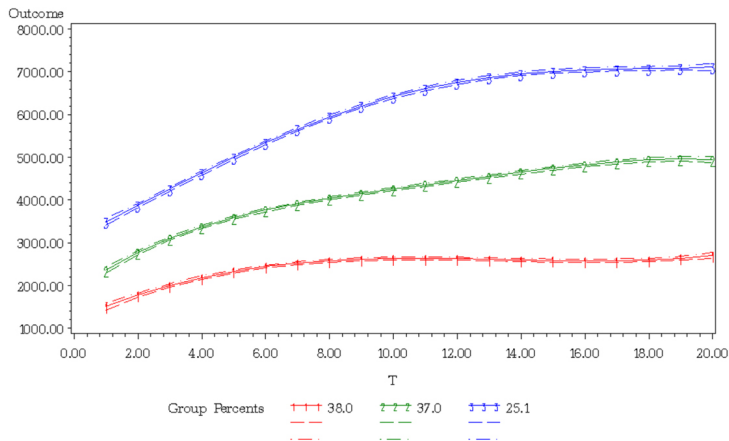
Result for 3 groups : workers beginning their career in 1984



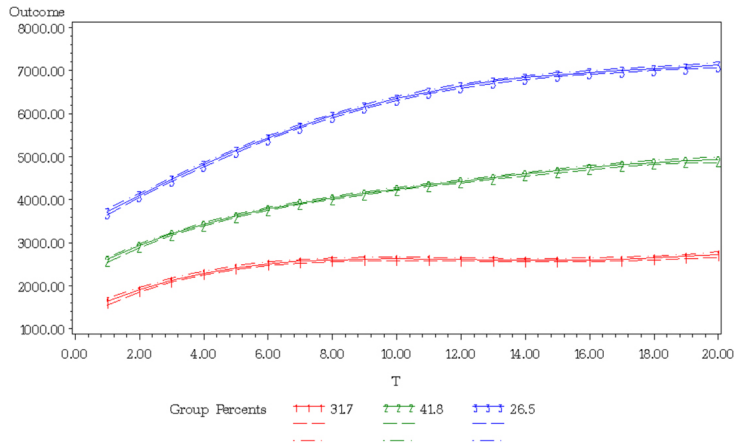
Result for 3 groups : workers beginning their career in 1985



Result for 3 groups : workers beginning their career in 1986



Result for 3 groups : workers beginning their career in 1987



Previous work

- Sampson R.J., Laub J.H. and Eggleston E.P. 2004. On the Robustness and Validity of Groups. *Journal of Quantitative Criminology* **20-1** p.37-42.
- Nagin D.S. and Tremblay R.E. 2005. Developmental trajectory groups: Fact or a useful statistical fiction? *Journal of Criminology* **43-4** p.873-904.
- Sampson R.J. and Laub J.H. 2005. Seductions of method: Rejoinder to Nagin and Tremblay's "Developmental trajectory groups: fact or fiction?". *Journal of Criminology* **43-4** p.905-913.

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Remark:

This approach is just useful to compare a whole set of models.

The mean shape

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If X denotes a random variable defined on a probability space $(\Omega, \mathcal{F}, \mathcal{P})$ with values in a metric space (Ξ, d) , an element $m \in \Xi$ is called a mean of $x_1, x_2, \dots, x_k \in \Xi$ if

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That means that the mean shape is defined as the shape with the smallest variance of all shapes in a group of objects.

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The test hypotheses are:

$$\text{Hypothesis: } H_0 : P = Q$$

$$\text{Alternative: } H_1 : P \neq Q$$

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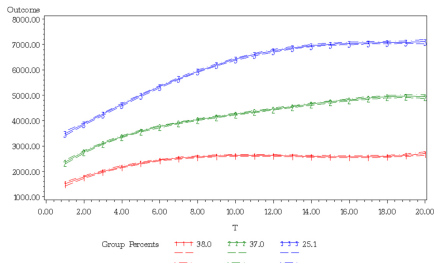
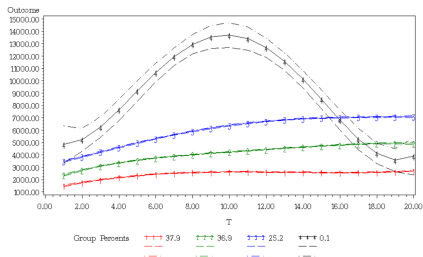
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- 5 Calculate the p -value for H_0 . $p_{r=i} = \frac{1}{\binom{N}{n}}$ for $i = 1, \dots, \binom{N}{n}$, where r is the rank for which we assume a uniform distribution.

The statistical shape analysis approach

Are these sets of trajectories different?

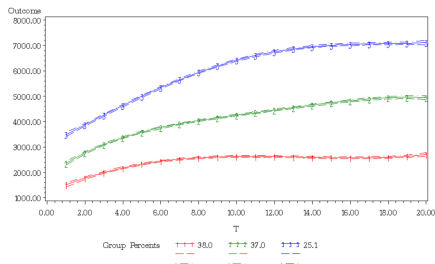
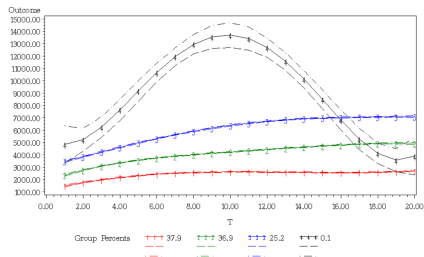
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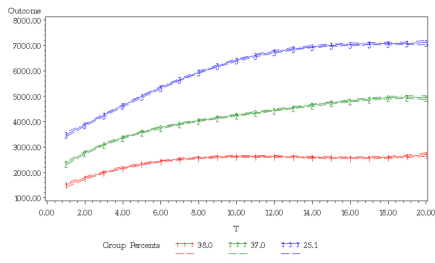
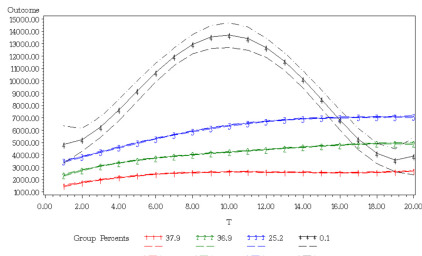
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Shape Analysis says yes,

The statistical shape analysis approach

Are these sets of trajectories different?



Shape Analysis says yes, but are they really?

The statistical shape analysis approach

Alternative methodology

To avoid this kind of situation, one can take the estimated parameters of the model as landmarks and perform a statistical "shape" analysis on these.

The classical statistics approach

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Compare the estimated parameters:

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Compare the estimated parameters:

- Performing the Wald test to see if the parameters differ between two models.
- Compare the confidence intervals of the parameters and see if they have an intersection.

Functional Data Analysis Approach

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Compare the set of trajectories as functions:

Functional Data Analysis Approach

Compare the set of trajectories as functions:

Consider a metrical space on the continuous functions defined on the time interval of the trajectories and use tests on functional data to analyze the time stability of the results.

Bibliography

- Nagin, D.S. 2005: *Group-based Modeling of Development*. Cambridge, MA.: Harvard University Press.
- Jones, B. and Nagin D.S. 2007: Advances in Group-based Trajectory Modeling and a SAS Procedure for Estimating Them. *Sociological Research and Methods*, **35** p.542-571.
- Guigou, J.D, Lovat, B. and Schiltz, J. 2012: Analysis of the salary trajectories in Luxembourg : a finite mixture model approach. To appear.
- Giebel S. 2011: *Zur Anwendung der statistischen Formanalyse*. Phd Thesis, University of Luxembourg.
- Schiltz, J. 2012: Robustness of groups and trajectories in Nagin's finite mixture model. To appear.