

FULL-LENGTH ORIGINAL RESEARCH

Association of ultra-rare coding variants with genetic generalized epilepsy: A case–control whole exome sequencing study

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Abstract

Objective: We aimed to identify genes associated with genetic generalized epilepsy (GGE) by combining large cohorts enriched with individuals with a positive family history. Secondly, we set out to compare the association of genes independently with familial and sporadic GGE.

Methods: We performed a case–control whole exome sequencing study in unrelated individuals of European descent diagnosed with GGE (previously recruited and sequenced through multiple international collaborations) and ancestry-matched controls. The association of ultra-rare variants (URVs; in 18 834 protein-coding genes) with epilepsy was examined in 1928 individuals with GGE (vs. 8578 controls), then separately in 945 individuals with familial GGE (vs. 8626 controls), and finally in 1005 individuals with sporadic GGE (vs. 8621 controls). We additionally examined the association of URVs with familial and sporadic GGE in two gene sets important for inhibitory signaling (19 genes encoding γ -aminobutyric acid type A [GABA_A] receptors, 113 genes representing the GABAergic pathway).

Results: *GABRG2* was associated with GGE ($p = 1.8 \times 10^{-5}$), approaching study-wide significance in familial GGE ($p = 3.0 \times 10^{-6}$), whereas no gene approached a significant association with sporadic GGE. Deleterious URVs in the most intolerant subgenomic regions in genes encoding GABA_A receptors were associated with familial GGE (odds ratio [OR] = 3.9, 95% confidence interval [CI] = 1.9–7.8, false

See Appendix 1 for list of contributors in Canadian Epilepsy Network, Epi4K Consortium, Epilepsy Phenome/Genome Project, EpiPGX Consortium, and EuroEPINOMICS-CoGIE Consortium. Affiliations are provided in the Supplementary Material.

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discovery rate [FDR]-adjusted $p = .0024$), whereas their association with sporadic GGE had marginally lower odds (OR = 3.1, 95% CI = 1.3–6.7, FDR-adjusted $p = .022$). URVs in GABAergic pathway genes were associated with familial GGE (OR = 1.8, 95% CI = 1.3–2.5, FDR-adjusted $p = .0024$) but not with sporadic GGE (OR = 1.3, 95% CI = .9–1.9, FDR-adjusted $p = .19$).

Significance: URVs in *GABRG2* are likely an important risk factor for familial GGE. The association of gene sets of GABAergic signaling with familial GGE is more prominent than with sporadic GGE.

KEYWORDS

familial epilepsy, GABA_A receptors, *GABRG2*, GGE, sporadic epilepsy

1 | INTRODUCTION

The genetic risk factors of generalized epilepsies have proven challenging to decipher despite evidence of their heritability from twin and family studies.^{1,2} Initial gene discovery, guided by linkage analysis, was performed in large families with autosomal dominant inheritance, but these cases proved rare³ and thus not necessarily representative of generalized epilepsies. Subsequently, both genome-wide association studies^{4,5} and rare variant association studies^{6–9} investigated increasingly larger cohorts of genetic generalized epilepsy (GGE). These studies provided key insights into the heritability and genetic architecture of GGE, which seems to involve ultra-rare genetic variants,^{6,8,9} common variants,^{4,5,10} and copy number alterations.^{11–13} Repeat expansions have also been recently implicated in dominantly inherited familial adult myoclonic epilepsy syndromes.^{14–17}

Prior large-scale sequencing studies of individuals with familial GGE failed to show statistically significant associations in single genes.^{6,7} Nonetheless, gene set burden analyses in these studies demonstrated that ultra-rare coding variants (URVs) in multiple phenotypically and biologically informed gene sets (e.g., dominant epilepsy and developmental epileptic and encephalopathy [DEE] genes, genes encoding γ -aminobutyric acid type A [GABA_A] receptors) are associated with an increased disease risk.^{6,7} These patterns were later replicated in independent case–control studies of predominantly sporadic GGE cases, which found few single genes approached study-wide significance despite much larger cohorts.^{8,9} A paradigm in which familial and sporadic epilepsy may have different genetic architectures has been previously established by our work on non-acquired focal epilepsy (NAFE) demonstrating a markedly higher burden of URVs in familial compared to sporadic NAFE.⁶ This, however, has not been investigated so far in GGE.

Key Points

- Although not significant study-wide, *GABRG2* is likely an important risk gene for GGE
- Compared to controls, URVs in *GABRG2* are seen more frequently in individuals with a familial GGE than a sporadic GGE
- Similarly, the association of URVs in GABAergic pathway genes is more prominent with familial GGE than with sporadic GGE

Aiming to identify protein-coding genes where URVs are significantly associated with an increased risk of generalized epilepsy, we performed a combined analysis of multiple cohorts of individuals with GGE and ancestry-matched controls. To improve the power of genetic discovery, we enriched our analysis with individuals with a positive family history of the disease, and also examined this subset of familial GGE separately. In additional analyses, we investigated individuals with sporadic GGE to determine whether familial and sporadic GGE have different genetic architectures.

2 | MATERIALS AND METHODS

2.1 | Study design and participants

In this case–control rare variant association study, we investigated the association of ultra-rare and rare genetic variants with epilepsy in individuals with a diagnosis of GGE and matched controls of European descent. We jointly analyzed whole exome sequencing (WES) data from two independent datasets encompassing GGE patients previously studied by (1) the Epi4K Consortium and the Epilepsy Phenome/Genome Project⁶ (referred

to hereafter, along with matched controls, as the first dataset) or (2) the Canadian Epilepsy Network and the EpiPGX and EuroEPINOMICS-CoGIE Consortia⁷ (referred to, with their matched controls, as the second dataset). Control cohorts were obtained for the first dataset from local collections available at the Institute for Genomic Medicine^{9,18} (IGM; New York, New York, USA), and for the second dataset from controls available at the Luxembourg Centre for Systems Biology (LCSB; Esch-sur-Alzette, Luxembourg) obtained from the database of Genotypes and Phenotypes¹⁹ or the Epi25 Collaborative.⁸ Ethical approvals from institutional review boards and relevant ethics committees and written informed consent procedures were previously obtained and detailed elsewhere.^{6,7} The details of the recruitment or acquisition of analyzed case or control cohorts, and diagnostic and inclusion criteria, were also previously described.^{6–8} Here, we intended primarily to identify genes significantly increasing the risk of GGE by combining these cohorts. To that aim, we analyzed data from 2203 affected individuals (1214 from the first dataset and 989 from the second dataset; before quality control). Subsequently, we examined the strength of the association separately in 1035 individuals (659 from the first dataset; 376 from the second dataset) with a positive family history of epilepsy. Afterward, we went on to assess the remaining 1168 individuals (555 from the first dataset; 613 from the second dataset) without a family history or with an unknown family history status.

2.2 | Sequencing and quality control

WES data generation for the case and control cohorts was previously described.^{6–8} In compliance with privacy regulations, the genotypes from the two datasets were processed in parallel at the IGM and the LCSB. A neural network predictive model was used to exclude individuals unlikely to be of non-Finnish European descent. We removed one sample from each pair of duplicates/related individuals within each dataset and one sample from each pair of duplicates between the two datasets. We also performed quality control procedures to remove low-quality samples/variants as well, to harmonize the coverage and call rate between the cases and controls within each dataset. Contingent on case–control matching, the final number of cases or controls included in each analysis (all, familial, and sporadic GGE analyses) differed slightly across analyses (see Results). The joint analysis strategy and the quality control procedures are outlined in Figure S1 and detailed in the supplemental methods (Supplementary Material).

2.3 | Variant annotations

The analysis was limited to coding variants located in the exons of 18 834 protein-coding genes from the Consensus Coding Sequence database (release 20),²⁰ extended with two bases on each side to accommodate canonical intronic splice sites. Variant effects were annotated using ClinEff (v1.0c).²¹ Population allele frequencies were estimated from the Genome Aggregation Database (gnomAD; r2.1)²² and DiscovEHR database (v1).²³ Because a portion of our control samples overlapped with gnomAD r2.1 exomes (see the Supplementary Material), gnomAD allele frequencies were based on gnomAD r2.1 genomes. Missense variants were further annotated with three *in silico* deleteriousness and intolerance scores (selected based on our previous work^{6,9}): PolyPhen-2 (PPh2) Human Diversity based score,²⁴ the Rare Exome Variant Ensemble Learner (REVEL) score,²⁵ and the Missense Tolerance Ratio (MTR) v1 score.²⁶ The population allele frequencies and *in silico* missense deleteriousness and intolerance scores were annotated for the first dataset (and its matched controls) using the Analysis Tool for Annotated Variants (ATAV) platform¹⁸ and for the second dataset (and its matched controls) using Annovar²⁷ and bcftools.²⁸

2.4 | Analysis models

We defined three primary analysis models to examine the association of functional coding variation with GGE, based on a combination of three filtering criteria: minor allele frequency, variant types (effects), and *in silico* predictions (specifically for missense variants). We targeted URVs, which we defined as those with a minor allele frequency (MAF) <.05% in our test datasets (internal MAF) and not seen in the independent gnomAD and DiscovEHR population reference datasets (external MAF). Functional variants (i.e., presumed to affect the function of protein-coding gene products) included those with predicted loss of function (pLoF; canonical splice-site, stop-gain/-loss, and frameshift variants), in-frame insertions and deletions, and missense variants. For each of the three models, missense variants were filtered further based on their expected (*in silico*) deleteriousness predicted using PPh2 or REVEL, or based on REVEL in combination MTR to capture the degree of subgenic intolerance at the affected site. The latter approaches based on REVEL and MTR (i.e., analysis of deleterious variants identified with an ensemble method designed for rare variants in combination with subgenic intolerance limiting) were recently shown to improve pathogenicity prediction in epilepsy and other disorders.^{9,26,29} A control model targeting synonymous URVs presumed to have a neutral effect was used to assess

potential biases in cases versus controls comparisons that are unlikely to be related to disease risk. We supplemented our primary analyses with additional secondary models to examine the association of (1) rare functional variants (defined as those with both internal and external MAFs < .1%) with and without URVs and (2) pLoF variants without other types of functional variants (as these represent a class of high-effect variants). Altogether, nine models were investigated (one control model, three primary models, and five secondary models) as summarized in Table 1.

2.5 | Gene-level associations

As adopted in our previous studies,⁶ we performed gene collapsing analyses by assigning a 1 or 0 indicator in a *gene by sample* matrix to indicate the presence or absence (respectively) of qualifying variants (QVs). QVs were defined as variants matching the criteria for each analysis model in a given gene and study individual (assuming dominant inheritance). The collapsing analysis was performed separately in our two independent study datasets, and a Cochran–Mantel–Haenszel (CMH) exact test was then used to quantify the gene-level association between case status and QV carrier status (by comparing the counts of cases and controls with QVs in the two datasets while accounting for cohort stratification).⁹ Separate comparisons were performed for all, familial, and sporadic GGE cohorts, each against their ancestry-matched controls. We adopted a Bonferroni multiple testing correction for gene-level *p*-values ($\alpha = .05$) accounting for three phenotypic groups, three primary analysis models, and 18 834 protein-coding genes with a study-wide significance cutoff of 2.9×10^{-7} . The homogeneity in the observed odds between the two datasets was examined using Breslow–Day and Woolf tests. The genomic inflation factor (λ) was estimated as detailed in the supplementary methods (see the Supplementary Material). The collapsing and subsequent joint statistical analyses were performed using ATAV¹⁸ or R data.table,³⁰ R tidyverse,³¹ and R stats³² on R v3.3.³²

2.6 | Gene set associations analyses

We also studied two gene sets that are important for inhibitory signaling – gene sets in which subjects with GGE had previously shown an increased burden of deleterious URVs. This association was established in a subset of our current samples⁷ and was later validated in additional datasets.⁸ However, a stratified analysis based on family history was not performed in our previous work. Here, we examined the association of URVs in these gene sets with familial GGE (vs. controls), with sporadic GGE (vs.

controls), and directly between individuals with familial GGE versus those with sporadic GGE. We complemented these comparisons with an analysis of all individuals with GGE versus controls (as a positive control). To measure the association, we did gene set collapsing analyses by collapsing QVs across all genes in the investigated gene set (i.e., a case/control was a carrier if they harbored a QV in any gene in the gene set) followed by CMH test. Probability values from the analyses of functional variants were adjusted for 24 multiple tests (four phenotypic comparisons, three URV analysis models, and two gene sets) using a Benjamini–Hochberg false discovery rate procedure to maximize the power (as opposed to Bonferroni correction for familywise error rate).

3 | RESULTS

We studied the association of coding URVs with generalized epilepsy in a cohort of 1928 unrelated individuals diagnosed with familial or sporadic GGE and 8578 matched controls of European descent. We also performed separate association studies for individuals diagnosed with familial GGE ($n = 945$, studied against 8626 matched controls) and individuals with a diagnosis of sporadic GGE ($n = 1005$, studied against 8621 matched controls; all counts after quality control). As case–control matching on principal components was performed separately for each analysis, the total number of controls differed slightly between the analyses, and the total number of samples in the analysis of all individuals with GGE was slightly less than the total of familial and sporadic cases. The sample counts from the two study cohorts are detailed in Table S1 (Supplementary Material). We did not detect a prominent deviation of observed *p*-values from expected *p*-values in synonymous collapsing analysis ($\lambda = .86$ – 1.06 ; Figure S5), indicating adequate population substructure matching between individuals with epilepsy (cases) and without epilepsy (controls).

No single gene achieved study-wide significance. However, *GABRG2* (Mendelian Inheritance in Man³³ [MIM] gene number: 137164) was the top-ranked gene in the analysis of the combined GGE cohort, showing prominent association in two primary models; it had a *p*-value of 1.8×10^{-5} in the PPh2 model (examining the association of functional URVs while filtering missense variants based on a damaging PPh2 prediction; Figure 1) and a *p*-value of 1.2×10^{-5} in the MTR model (combining subgenic intolerance filtering with REVEL; Figure S7). Limiting the cases to individuals with a family history of epilepsy strengthened the association with *GABRG2* in the PPh2 model ($p = 3.0 \times 10^{-6}$). Using REVEL combined with MTR did not outperform the PPh2 model in terms of

TABLE 1 Overview of association analysis models

	Models									
	Primary models					Secondary models				
	Ultra-rare functional variants					Rare functional variants				
Control	Synonymous	PPh2	REVEL	MTR	Rare variants + URVs	Rare variants - URVs	Rare variants + URVs	Rare variants - URVs	Rare variants + URVs	Rare variants - URVs
MAFs										
Internal MAF	<.0005	<.0005	<.0005	<.0005	<.0005	<.001	Without URVs	<.001	Without URVs	<.0005
DiscovEHR MAF	0	0	0	0	0	<.001	URVs	<.001	URVs	0
gnomAD r2 MAF	0	0	0	0	0	<.001	<.001	<.001	<.001	0
Classes of variants										
ClinEff effects & missense variants filters	Synonymous	Functional	Functional	Functional	Functional	Functional	Functional	Functional	pLoF	pLoF
PPh2 prediction	-	"Probably"	-	-	-	"Probably"	"Probably"	"Probably"	-	-
REVEL score	-	-	≥.5	≥.5	≥.5	-	-	-	-	-
MTR score	-	-	-	≤.78	≤.78	-	-	-	-	-

Note: pLoF variants included stop-gain and stop-loss variants, frameshift insertions and deletions, and canonical splice-site variants. Functional variants included pLoF, in-frame insertions and deletions, and missense variants (the missense variants were filtered using PPh2, REVEL, and MTR predictions as indicated). MAFs from gnomAD were based on the "Genomes" subset. The cutoffs for REVEL and MTR scores were based on the previous study of the Epi25 Collaborative.⁸

Abbreviations: MAF, minor allele frequency; MTR, Missense Tolerance Ratio score; pLoF, predicted loss-of-function variants; PPh2, PolyPhen-2 Human Diversity-based prediction; REVEL, Rare Exome Variant Ensemble Learner; URV, ultra-rare coding variant.

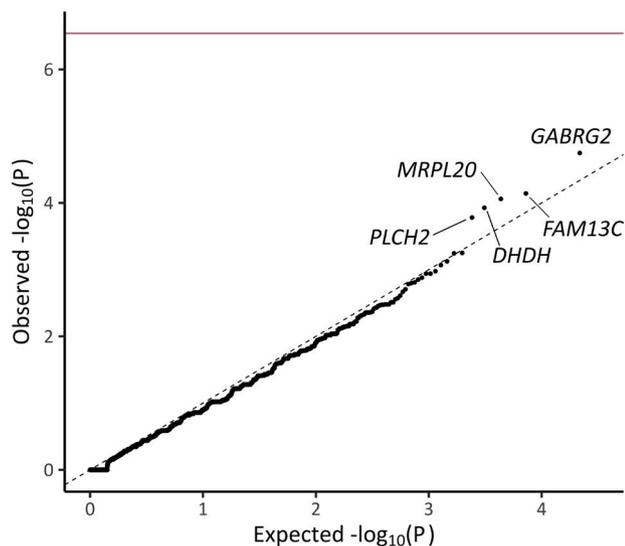
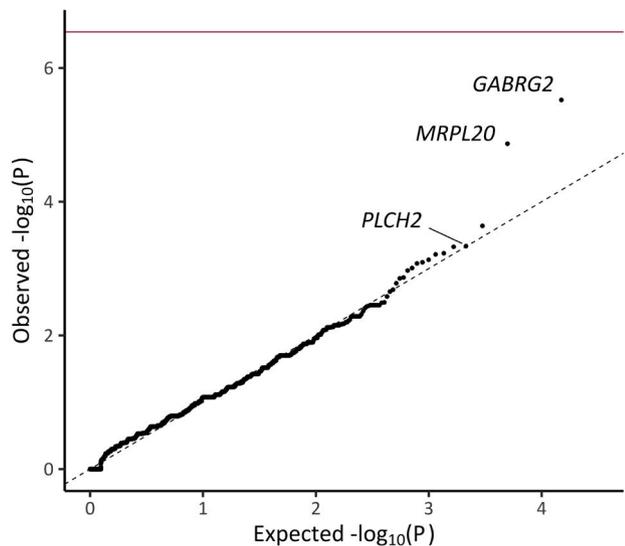
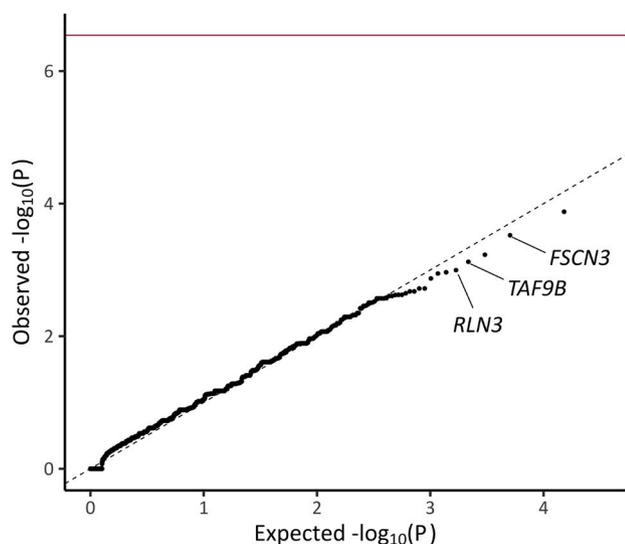
Functional URVs (PPh2_{Damaging})All GGEs. λ : 0.91Familial GGEs. λ : 1.03Sporadic GGEs. λ : 1.05

FIGURE 1 Association of ultra-rare variation in protein-coding genes with genetic generalized epilepsy. The quantile–quantile plots compare the observed p -values (Cochran–Mantel–Haenszel exact test) and the expected p -values (drawn from a uniform distribution) in analyses of 1928 subjects with genetic generalized epilepsy (GGE) in comparison to 8578 matched controls (top panel) as well as subsequent analyses of subjects with familial GGE (middle panel; 945 cases and 8626 controls) and subjects with sporadic GGE (bottom panel; 1005 cases and 8621 controls). These analyses focused on functional ultra-rare variants (URVs; with minor allele frequencies $< .05\%$ in the test dataset, and not seen in DiscovEHR/gnomAD) that were annotated as predicted loss-of-function variants (pLoF), damaging missense variants as predicted with PolyPhen-2 (PPh2), or in-frame insertions and deletions. Study-wide significance after Bonferroni correction (dark red line) was defined by a p -value of $< 2.9 \times 10^{-7}$. λ , genomic inflation factor. Among the five top-ranked genes, genes that had a higher carrier frequency in cases versus controls in both study datasets are labeled (not labeled among top-ranked: enriched in controls or in cases in one dataset only)

significance in the analysis of subjects with familial GGE ($p = 1.4 \times 10^{-5}$). Nonetheless, it maximized the separation between cases and controls, resulting in higher odds by preferentially filtering all *GABRG2* variants seen in our control sets (Tables 2, S2, and S3).

The analysis of subjects with sporadic GGE was generally unremarkable for *GABRG2* ($p = .15$ – $.015$), and the top-ranked genes did not include biologically meaningful candidates (Table 2). Rare variant analyses (up to an MAF of $.1\%$) resulted in the inclusion of additional *GABRG2* variants exclusively in the control cohorts (Table S3). In general, secondary analyses of rare functional and pLoF variants captured neither significantly associated single genes nor strong novel candidates with biological relevance (Tables S4 and S5; Figures S8 and S9). Although not study-wide significant, *GABRG2* achieved a higher rank than in our prior URV analysis⁶ in 640 subjects with familial GGE versus 3877 controls using an analysis model comparable to the current PPh2 model (Rank 7, $p = 9.2 \times 10^{-4}$). Its rank was higher than that seen in two recent large-scale analyses from the Epi25 Collaborative^{8,9} in 3108 individuals with GGE versus 8436 controls (Rank 3, $p = 6.2 \times 10^{-4}$) and 5303 subjects with GGE versus 15 677 controls (Rank 37, $p = 6.1 \times 10^{-3}$).

Most URVs in *GABRG2* were missense, and four URVs were seen in individuals with absence seizures (Table S2). Eight variants were seen in the familial GGE cohort, including two that were confirmed to be inherited; p.R177P, identified in a proband with early onset absence epilepsy, was inherited from a parent with a similar phenotype, whereas p.Y213* was inherited from a parent not diagnosed with epilepsy (Figure S10). We did not have

TABLE 2 Top-ranked genes in the primary analyses of ultra-rare functional variants

Analysis	URVs	HGNC	Epilepsy gene	Qualifying cases			Qualifying controls			OR (95% CI)	p (homogeneity)
				1st dataset	2nd dataset	Both datasets	1st dataset	2nd dataset	Both datasets		
All GGE	PPh2	GABRG2	Yes	7 (.64%)	3 (.36%)	10 (.52%)	3 (.04%)	1 (.06%)	4 (.05%)	12.1 (3.4–54.1)	1.8×10^{-5} (.54)
	REVEL			4 (.36%)	3 (.36%)	7 (.36%)	0 (.00%)	1 (.06%)	1 (.01%)	28.3 (3.4–1307.3)	1.3×10^{-4} (.15)
	MTR			4 (.36%)	3 (.36%)	7 (.36%)	0 (.00%)	0 (.00%)	0 (.00%)	∞ (6.1– ∞)	1.2×10^{-5} (.53)
Familial GGE	PPh2	GABRG2	Yes	6 (.95%)	2 (.63%)	8 (.85%)	3 (.04%)	1 (.06%)	4 (.05%)	18.9 (5–86.5)	3.0×10^{-6} (.63)
	REVEL			3 (.48%)	2 (.63%)	5 (.53%)	0 (.00%)	1 (.06%)	1 (.01%)	40.6 (4.4–1934.3)	1.0×10^{-4} (.19)
	MTR			3 (.48%)	2 (.63%)	5 (.53%)	0 (.00%)	0 (.00%)	0 (.00%)	∞ (7.9– ∞)	1.4×10^{-5} (.64)
Sporadic GGE	PPh2	FAM13C	-	5 (.81%)	0 (.00%)	5 (.50%)	4 (.05%)	0 (.00%)	4 (.05%)	17.6 (3.8–89.0)	1.3×10^{-4} (.44)
	REVEL	TNFRSF21	-	2 (.47%)	2 (.39%)	4 (.40%)	0 (.00%)	0 (.00%)	0 (.00%)	∞ (5.1– ∞)	2.3×10^{-4} (.52)
	MTR	TRPV5	-	3 (.70%)	0 (.00%)	3 (.30%)	0 (.00%)	0 (.00%)	0 (.00%)	∞ (5.8– ∞)	3.0×10^{-4} (.18)

Note: ORs and p-values are given from a Cochran–Mantel–Haenszel exact test. No gene reached study-wide significance ($p < 2.9 \times 10^{-7}$). The accompanying homogeneity p-value indicates the lowest p-value from Breslow-Day and Woolf tests for homogeneity of odds, where p-values < .05 indicate significantly different odds between the two analysis datasets. See Table 1 for the details of the PPh2, REVEL, and MTR analysis models.

Abbreviations: CI, confidence interval; GGE, genetic generalized epilepsy; HGNC: HUGO [Human Genome Organization] Gene Nomenclature Committee gene symbols. MTR, Missense Tolerance Ratio; OR, odds ratio; PPh2, PolyPhen-2; REVEL, Rare Exome Variant Ensemble Learner; URV, ultra-rare coding variant.

sufficient data and samples to determine the allelic origin of the remaining variants (p.A160S, p.M199V, p.V252A, p.G413S, p.D450Y, and p.N456S). Two other canonical splice donor variants (IVS2+2G>T and IVS6+2G>T) were seen in our sporadic GGE cohort. The predicted protein changes for these variants are based on the transcript NM_198904.

Few of these URVs were recurrent. p.M199V (familial GGE) was reported in a previous study,³⁴ in which it segregated with a phenotype of generalized epilepsy with febrile seizures plus (GEFS+) and also in an individual with NAFE in the first Epi25 Collaborative study.⁸ p.R177P (familial GGE) affected a codon for which a different change (p.R177G) was seen previously in a family with febrile seizures (FS).³⁵ IVS2SD (sporadic GGE) was previously reported in ClinVar (ID: VCV001067627.1) as likely pathogenic for childhood absence epilepsy (CAE) and FS (no details on family history). Last, IV6SD was previously associated with familial CAE and FS.³⁶ The patient included here (sporadic GGE) had absence epilepsy but no history of FS. Sample overlap or relatedness to these previously reported individuals was not investigated genetically, but it was considered unlikely based on our patients' clinical and family histories.

Apart from *GABRG2*, there was little overlap between the leading associations in the recent analyses^{6,8,9} and this study (Table S6). *CACNA1B* (MIM: 601012), the top hit in our prior analysis ($p = 1.7 \times 10^{-5}$),⁶ showed a less prominent association than previously seen (Rank 5 in the MTR model/familial GGE; $p = .00098$). Our analysis also did not recapture two genes previously seen as top hits with suggestive association (*CACNA1G* [MIM: 604065] with $p = 2.5 \times 10^{-4}$ and *SLC6A1* [MIM: 137165] with $p = 2.1 \times 10^{-6}$).^{8,9} *GABRA1* (MIM: 137160) was among a few shared top hits, achieving comparable ranks in all studies (Rank 9 in the MTR model analysis of all subjects with GGE with $p = .0023$; Rank 8 in the first Epi25 Collaborative study⁸ with $p = .0022$; Rank 9 in the second study⁹ with $p = .0013$). Few other MIM genes previously suggested to increase the susceptibility to GGE were among the top hits (Tables S7 and S8). On the other hand, several MIM genes underlying dominant DEEs were among the top-ranked genes (Table S9), as expected from the known enrichment of such URVs in genes causing dominant DEEs in generalized epilepsies.^{6,8,9}

The association of URVs in two gene sets important for GABAergic signaling (genes encoding GABA_A receptors and GABAergic pathway genes) with the phenotype was not prominent in the analysis of deleterious URVs, whereas the incorporation of subgenic intolerance in the definition of QVs improved the power⁹ and unraveled clear association signals in the analysis of all subjects with GGE (Figure 2A) and subjects with familial GGE (Figure 2B).

It also aided the identification of an association between genes encoding GABA_A receptors and sporadic GGE, although weaker than that observed in comparisons of subjects with familial GGE versus controls (Figure 2C). We did not detect an association between GABAergic pathway genes and sporadic GGE as expected from previous findings,⁸ possibly due to insufficient power or differences in the analysis models (Figure 2C). The outcomes of a direct comparison of 945 individuals with familial GGE versus 1005 individuals with sporadic GGE were unremarkable and also likely underpowered (Figure 2D).

4 | DISCUSSION

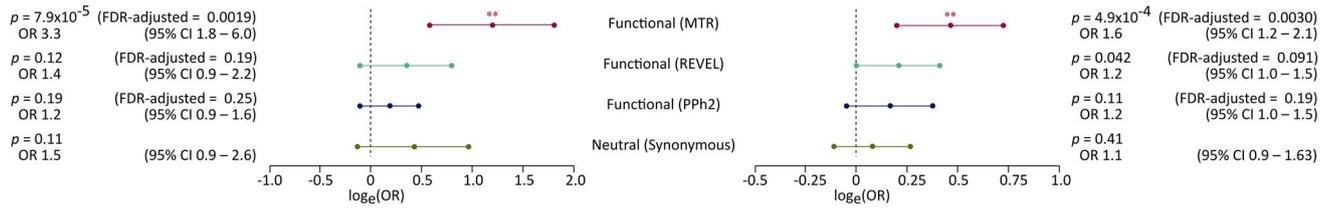
Here, we add to the evidence indicating that deleterious URVs in *GABRG2* are a risk factor for generalized epilepsies, although this gene did not reach study-wide significance. Notably, this association appears to be driven by ultra-rare private variants rather than rare variants (possibly seen in external population controls). This work emphasizes the role of ultra-rare variation in less severe epilepsies and corroborates the association of coding variation in *GABRG2* with familial GGE.^{6,8,9,37} The current analysis benefits from a higher number of individuals with familial GGE and a balanced distribution of familial and sporadic cases compared to recent large-scale analyses^{8,9} enriched for subjects with sporadic GGE. Our attempts to integrate multiple cohorts from independent studies to achieve this larger sample size came with some limitations. Quality control and harmonization measures mandated the exclusion of putative QVs in genes of interest. The restrictions in genotype sharing across study sites limited the possibilities to invoke analysis methods incorporating covariates to handle residual population stratification. Also, the use of phenotypic definitions and classifications from independent studies might have resulted inadvertently in minor inconsistencies in sample stratification across the familial and sporadic cohorts (which included individuals with unknown family history status).

Absence seizures, a seizure type that was prominent in earlier *GABRG2* families featuring an overlap of GGE and GEFS+,³⁸⁻⁴⁰ were also predominant among individuals with QVs in *GABRG2* in this cohort (Table S2). The phenotypes in individuals with a positive family history and their affected siblings or parents were mostly congruent (Figure S10). The small number of affected individuals and the limited segregation analysis precluded reliable estimation of penetrance or heterogeneity. Although the segregation of *GABRG2* variants could be studied in only two families, these showed that pathogenic variants could be inherited from both affected and nonaffected parents.

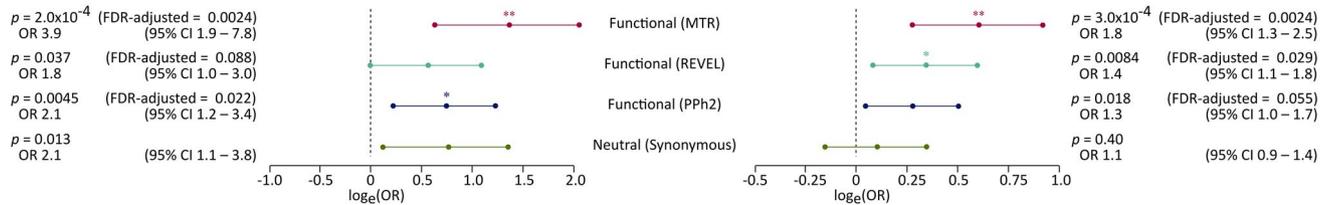
Genes encoding GABA_A receptors subunits

Genes representing the GABAergic pathway

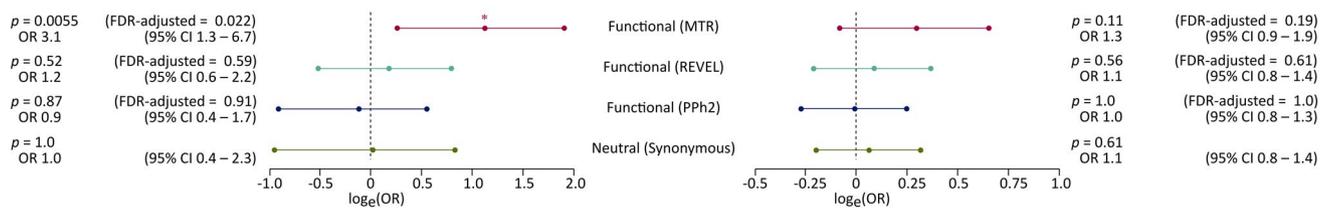
(A) Genetic generalized epilepsy vs. controls



(B) Familial genetic generalized epilepsy vs. controls



(C) Sporadic genetic generalized epilepsy vs. controls



(D) Familial vs. sporadic genetic generalized epilepsy

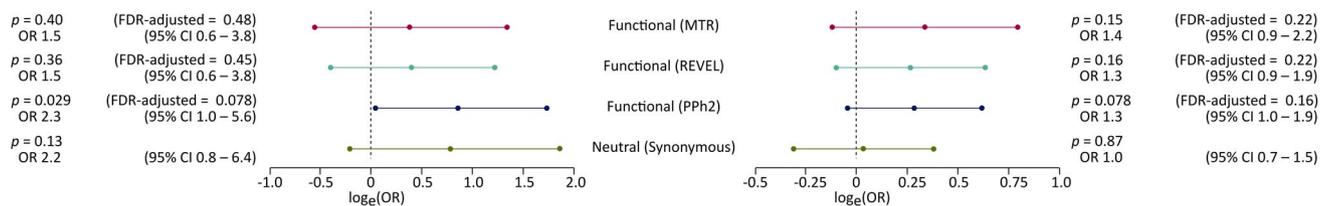


FIGURE 2 Association of ultra-rare variation in genes encoding γ -aminobutyric acid type A (GABA_A) receptors with familial and sporadic genetic generalized epilepsy (GGE). The forest plots show the association of ultra-rare deleterious and intolerant variants with the phenotype in analyses of 1928 individuals with GGE versus 8578 controls (A), 945 individuals with familial GGE versus 8626 controls (B), 1005 individuals with sporadic GGE versus 8621 controls (C), and a direct comparison of 945 individuals with familial GGE versus 1005 individuals with sporadic GGE (D). Four (primary and control) ultra-rare variant models are shown (y axis). The association in each analysis is displayed as the natural logarithm of stratified odds ratio (OR) from a Cochran–Mantel–Haenszel exact test (x axis). Errors bars indicate the logarithm of the 95% confidence intervals (CIs). The corresponding ORs and associated *p*-values, and false discovery rate (FDR)-adjusted *p*-values, are displayed on the side. Noteworthy FDR-adjusted *p*-values are indicated with stars (** < .005, * < .05). The tests for synonymous variants were not adjusted for multiple testing. MTR, Missense Tolerance Ratio; PPh2, PolyPhen-2; REVEL, Rare Exome Variant Ensemble Learner

This is concordant with prior observations that penetrance was typically incomplete and that *GABRG2*-related GGE had complex inheritance; most inherited pathogenic *GABRG2* variants had reduced penetrance, sometimes with phenotypic heterogeneity, whereas de novo variants were more prevalent in individuals with severe or developmental phenotypes (Table S10).

The lack of study-wide significance in rare variant association studies in GGE and the failure to reproduce multiple leading associations speak to the marked genetic heterogeneity in subjects with GGE. The exact extent of

the contribution of rare coding variation in GGE heritability is largely unknown. It remains, therefore, difficult to speculate on the interpretation of any negative findings, and on whether a further increase in statistical power might result in suggestive associations reaching significance. Using a similar study design to the one used to examine the current set (slightly exceeding 10 000 samples), we estimate that a total sample size exceeding 16 000 samples would be required to achieve study-wide significance in a gene with rates of QVs similar to those observed in *GABRG2*. These carrier rates seem, however, to be an

upper-bound estimate due to the multitude of subjects with familial GGE included here; the sample size required is probably much larger when examining subjects with sporadic GGE.^{8,9}

Nonetheless, the observed association of *GABRG2* with GGE further validates the outcomes of an analysis performed by the Epi25 Collaborative (albeit with partial overlap in datasets; see the Supplementary Material).⁸ The prominent difference in *GABRG2* rank in a second iteration⁹ of the Epi25 study with an expanded sample size might be explained by the familial origin of *GABRG2* variants; both studies had considerably lower ratios of familial to sporadic GGE (approximately a 1:7 ratio). *GABRG2* was also the lead association in a burden analysis of pLoF URVs ($p = 6.9 \times 10^{-5}$) in a recent study investigating the exomes of 3999 individuals with epilepsy (without further phenotypic subclassification) versus 277 586 controls from the UK Biobank.⁴¹ The different definitions of QVs in these studies might also explain the variable outcomes. Our PPh2 analysis model is similar to the prior model we used to analyze a subset of our samples (thus allowing for comparisons of outcomes with the increase in sample size).⁶

Compared to PPh2 filtering, *GABRG2* URVs had higher odds of association with GGE when missense variants were filtered using REVEL, in line with REVEL's higher performance in discriminating pathogenic and benign rare variants.²⁵ Additional filtering on subgenic intolerance (MTR) increased the odds further, consistent with recent findings suggesting that subgenic intolerance filtering is particularly effective for analyses geared toward specificity as opposed to sensitivity.^{9,29} REVEL and MTR cutoffs similar to those utilized in the most extensive and recent rare variant association study on epilepsy were used. Different values for these filters maximize the separation of benign and pathogenic variants in different types of epilepsy.⁹ However, most functionally validated *GABRG2* variants previously implicated in epilepsy fit one or more of the QV models we used, indicating good recall of disease-related variants with the current parameters (Table S10).

The recurrence of the same *GABRG2* variants in individuals with different types of epilepsy (GEFS+, DEE, GGE, NAFE; Table S10), as well as in familial and sporadic GGE with overlapping phenotypes, underscores a considerable genetic overlap and possibly a complex inheritance. Although we found the most substantial contribution from deleterious variants not seen in the gnomAD and DiscovEHR databases, a small contribution from rare variants or variants with benign predictions to this complex genetic predisposition cannot be ruled out (for instance, p.N79S previously identified in individuals with GGE or NAFE causes subtle functional alterations^{8,42-44}

and is seen in three individuals in gnomAD release 3.1.2). The *GABRG2* locus was recently found to be associated with FS,⁴⁵ highlighting the role of common variants in a phenotype that was prominent in earlier families with an increased susceptibility to GGE and GEFS+ linked to rare *GABRG2* variants.³⁸⁻⁴⁰

Prior burden analyses also revealed the presence of shared patterns of risk determinants between severe epilepsies (DEE) and common epilepsies (GGE, NAFE) in gene sets that are key for inhibitory signaling.^{7,8} A former analysis (in 3108 individuals with GGE) did not capture a considerable change in URV burden in genes encoding GABA_A receptors or GABAergic pathway genes upon the exclusion of a relatively small subset ($n = 380$) of familial samples.⁸ Conversely, we found a more prominent association between ultra-rare coding variation in GABAergic pathway genes and familial GGE in comparison to its association with sporadic GGE, albeit not demonstrable in direct (familial vs. sporadic) comparisons. Direct comparisons with sufficient power could help confirm the subtle differences in risk profiles.

In summary, we show that URVs in *GABRG2* are potentially an important risk factor for GGE, although not reaching study-wide significance. The association of URVs in genes representing the GABAergic pathway is likely more prominent in familial GGE than in sporadic GGE. Future work on epilepsy cohorts enriched with familial cases, extending the analysis to additional types of genetic variation (e.g., alterations in copy numbers and repeats, rare intronic and regulatory variants, and common risk alleles), could further our understanding of the genetic heterogeneity in GGE and the evidently complex inheritance.

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CONFLICT OF INTEREST

R.S.D. is a paid consultant of AstraZeneca. None of the other authors has any conflict of interest related to this article to disclose. We confirm that we have read the Journal's position on issues involved in ethical publication and affirm that this report is consistent with those guidelines. The funding agencies had no role in study design; in the collection, analysis, and interpretation of data; and in the writing and the decision to submit the paper for publication.

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Conception, planning, and design: Samuel F. Berkovic, Holger Lerche, David B. Goldstein, Daniel H. Lowenstein, P.M., K.E.S., R.S.D. Patients' and sequence data: CENet, Epi4K Consortium, EP/GP, EpiPGX Consortium, EuroEPINOMICS-CoGIE Consortium. Data analysis: M.K., J.E.M., K.E.S., D.R.B., R.S.D., P.M. Data interpretation: M.K., J.E.M., K.E.S., R.S.D., P.M., H.L., D.B.G., S.F.B., D.H.L. Writing—first draft: M.K. Writing—revisions: M.K., J.E.M., K.E.S., P.M., H.L., S.F.B., D.H.L., D.B.G.

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SUPPORTING INFORMATION

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APPENDIX 1

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