A Finite Mixture Model with Trajectories Depending on Covariates

Jang SCHILTZ (University of Luxembourg)

joint work with Jean-Daniel GUIGOU (University of Luxembourg), & Bruno LOVAT (University of Lorraine)

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1 The Basic Finite Mixture Model of Nagin



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Outline

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2 Generalizations of the basic model





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mixture : population composed of a mixture of unobserved groups



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Hence, this model can be interpreted as functional fuzzy cluster analysis.

Finite mixture model (Daniel S. Nagin (Carnegie Mellon University))

- mixture : population composed of a mixture of unobserved groups
- finite : sums across a finite number of groups



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where $P^{j}(Y_{i})$ is probability of Y_{i} if subject *i* belongs to group *j*.



<u>Aim of the analysis</u>: Find *r* groups of trajectories of a given kind (for instance polynomials of degree 4, $P(t) = \beta_0 + \beta_1 t + \beta_2 t^2 + \beta_3 t^3 + \beta_4 t^4$.



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We try to estimate a set of parameters $\Omega = \left\{ \beta_0^j, \beta_1^j, \beta_2^j, \beta_3^j, \beta_4^j, \pi_j \right\}$ which allow to maximize the probability of the measured data.



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Possible data distributions:

- count data \Rightarrow Poisson distribution
- binary data \Rightarrow Binary logit distribution
- \bullet censored data \Rightarrow Censored normal distribution



Notations :



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$$\beta^j t_{i_t} = \beta_0^j + \beta_1^j Age_{i_t} + \beta_2^j Age_{i_t}^2 + \beta_3^j Age_{i_t}^3 + \beta_4^j Age_{i_t}^4$$



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• $\phi :$ density of standard centered normal law.

Then,

$$L = \frac{1}{\sigma} \prod_{i=1}^{N} \sum_{j=1}^{r} \pi_j \prod_{t=1}^{T} \phi\left(\frac{y_{it} - \beta^j t_{it}}{\sigma}\right).$$



(2)

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It is too complicated to get closed-forms equations.





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SAS-based Proc Traj procedure

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Just implements a zero-inflation Poission model.



Model Selection (1)

Bayesian Information Criterion:

$$BIC = \log(L) - 0,5k\log(N), \qquad (3)$$

where k denotes the number of parameters in the model.

Rule:

The bigger the BIC, the better the model!


Model Selection (2)

Leave-one-out Cross-Validation Apporach:

$$CVE = \frac{1}{N} \sum_{i=1}^{N} \frac{1}{T} \sum_{t=1}^{T} \left| y_{i_t} - \hat{y}_{i_t}^{[-i]} \right|.$$
(4)

Rule:

The smaller the CVE, the better the model!



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Posterior Group-Membership Probabilities

Posterior probability of individual *i*'s membership in group $j : P(j/Y_i)$.

Bayes's theorem

$$\Rightarrow P(j/Y_i) = \frac{P(Y_i/j)\hat{\pi}_j}{\sum_{j=1}^r P(Y_i/j)\hat{\pi}_j}.$$
(5)

Bigger groups have on average larger probability estimates.

To be classified into a small group, an individual really needs to be strongly consistent with it.





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About 7 million salary lines corresponding to 718.054 workers.



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Some sociological variables:

• gender (male, female)



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- gender (male, female)
- nationality and residentship (luxemburgish residents, foreign residents, foreign non residents)



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- gender (male, female)
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- year of birth
- age in the first year of professional activity



Result for 9 groups (dataset 1)



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Result for 9 groups (dataset 1)





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Results for 9 groups (dataset 1)

Maximum Likelihood Estimates Model: Censored Normal (CNORM)

			Standard	T for HO:	
Group	Parameter	Estimate	Error	Parameter=0	Prob > T
1	Intercept	589.03067	18.46813	31.894	0.0000
	Linear	387.72145	11.31617	34.263	0.0000
	Quadratic	-14.36621	2.12997	-6.745	0.0000
	Cubic	-0.01563	0.15109	-0.103	0.9176
	Quartic	0.00856	0.00358	2.395	0.0166
2	Intercept	784.79156	15.75939	49.798	0.0000
	Linear	277.63602	9.78078	28.386	0.0000
	Quadratic	-28.36731	1.83236	-15.481	0.0000
	Cubic	1.17739	0.12972	9.076	0.0000
	Quartic	-0.01635	0.00307	-5.330	0.0000
3	Intercept	709.28728	15.90545	44.594	0.0000
	Linear	318.88029	8.97949	35.512	0.0000
	Quadratic	-21.54540	1.69611	-12.703	0.0000
	Cubic	0.62010	0.12002	5.167	0.0000
	Quartic	-0.00440	0.00284	-1.554	0.1203



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 x_i : vector of variables potentially associated with group membership (measured before t_1).



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Multinomial logit model:

$$\pi_j(x_i) = \frac{e^{x_i \theta_j}}{\sum\limits_{k=1}^r e^{x_i \theta_k}},$$
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where θ_i denotes the effect of x_i on the probability of group membership.



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$$L = \frac{1}{\sigma} \prod_{i=1}^{N} \sum_{j=1}^{r} \frac{e^{x_i \theta_j}}{\sum_{k=1}^{r} e^{x_i \theta_k}} \prod_{t=1}^{T} \phi\left(\frac{y_{i_t} - \beta^j t_{i_t}}{\sigma}\right).$$
(7)



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Let $z_1...z_L$ be covariates potentially influencing Y.



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We are then looking for trajectories

 $y_{i_{t}} = \beta_{0}^{j} + \beta_{1}^{j} Age_{i_{t}} + \beta_{2}^{j} Age_{i_{t}}^{2} + \beta_{3}^{j} Age_{i_{t}}^{3} + \beta_{4}^{j} Age_{i_{t}}^{4} + \alpha_{1}^{j} z_{1} + \dots + \alpha_{L}^{j} z_{L} + \varepsilon_{i_{t}},$ (8)

where $\varepsilon_{i_t} \sim \mathcal{N}(0, \sigma)$, σ being a constant standard deviation and z_l are covariates that may depend or not upon time t.



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where $\varepsilon_{i_t} \sim \mathcal{N}(0, \sigma)$, σ being a constant standard deviation and z_l are covariates that may depend or not upon time t.

Unfortunately the influence of the covariates in this model is limited to the intercept of the trajectory.





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Our model



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Our model

Let $x_1...x_L$ and $z_{i_1},...,z_{i_T}$ be covariates potentially influencing Y. We propose the following model:

$$\begin{aligned} y_{it} &= \left(\beta_0^j + \sum_{l=1}^L \alpha_{0l}^j x_l + \gamma_0^j z_{it}\right) + \left(\beta_1^j + \sum_{l=1}^L \alpha_{1l}^j x_l + \gamma_1^j z_{it}\right) Age_{it} \\ &+ \left(\beta_2^j + \sum_{l=1}^L \alpha_{2l}^j x_l + \gamma_2^j z_{it}\right) Age_{it}^2 + \left(\beta_3^j + \sum_{l=1}^L \alpha_{3l}^j x_l + \gamma_3^j z_{it}\right) Age_{it}^3 \\ &+ \left(\beta_4^j + \sum_{l=1}^L \alpha_{4l}^j x_l + \gamma_4^j z_{it}\right) Age_{it}^4 + \varepsilon_{it}, \end{aligned}$$

where $\varepsilon_{i_t} \sim \mathcal{N}(0, \sigma)$, σ being a constant standard deviation.



Men versus women



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Statistical Properties

Both Nagin's and our model can be written as

$$L = \frac{1}{\sigma} \prod_{i=1}^{N} \sum_{j=1}^{r} \pi_{j} \prod_{t=1}^{T} \phi\left(\frac{\text{observed data - modelled data}}{\sigma}\right). \tag{9}$$



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(9)

Nagin's model:

$$y_{i_{t}} = \beta_{0}^{j} + \beta_{1}^{j} Age_{i_{t}} + \beta_{2}^{j} Age_{i_{t}}^{2} + \beta_{3}^{j} Age_{i_{t}}^{3} + \beta_{4}^{j} Age_{i_{t}}^{4} + \alpha_{1}^{j} z_{1} + \dots + \alpha_{L}^{j} z_{L} + \varepsilon_{i_{t}},$$
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Our model:

$$y_{i_{t}} = \left(\beta_{0}^{j} + \sum_{l=1}^{L} \alpha_{0l}^{j} x_{l} + \gamma_{0}^{j} z_{i_{t}}\right) + \left(\beta_{1}^{j} + \sum_{l=1}^{L} \alpha_{1l}^{j} x_{l} + \gamma_{1}^{j} z_{i_{t}}\right) Age_{i_{t}}$$
$$+ \left(\beta_{2}^{j} + \sum_{l=1}^{L} \alpha_{2l}^{j} x_{l} + \gamma_{2}^{j} z_{i_{t}}\right) Age_{i_{t}}^{2} + \left(\beta_{3}^{j} + \sum_{l=1}^{L} \alpha_{3l}^{j} x_{l} + \gamma_{3}^{j} z_{i_{t}}\right) Age_{i_{t}}^{3}$$
$$+ \left(\beta_{4}^{j} + \sum_{l=1}^{L} \alpha_{4l}^{j} x_{l} + \gamma_{4}^{j} z_{i_{t}}\right) Age_{i_{t}}^{4} + \varepsilon_{i_{t}},$$

A way of estimating our model with the existing software (if group membership does not depend on the covariates) :



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• Use the latest version of proc.traj to test if the covariates have indeed an influence on the trajectories.



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A way of estimating our model with the existing software (if group membership does not depend on the covariates) :

- Use the latest version of proc.traj to test if the covariates have indeed an influence on the trajectories.
- Apply proc.traj to the data without covariates do the clustering and obtain the number of groups and the constitution of the groups.
- Use your favorite regression model software to get the trajectories separately for each group.



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